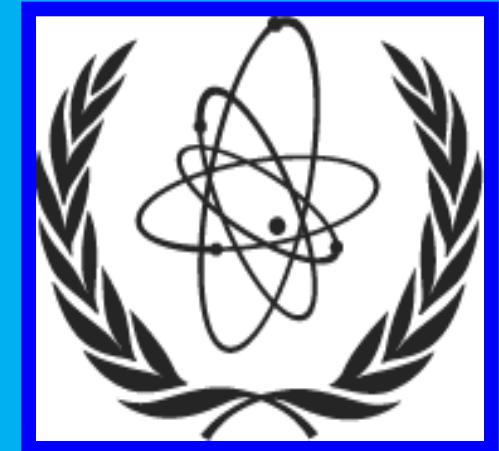


Unified Monte Carlo evaluation method



Roberto Capote and Andrej Trkov
IAEA Nuclear Data Section, Vienna, Austria
Donald L. Smith, Argonne National Laboratory, USA

(Nuclear) Data Evaluation

Evaluated cross sections and covariance matrices

Experimental Input

Inter and -intra
experiment
correlations

Experimental
cross sections

Prior Knowledge

Model Defects

Parameter
Uncertainties

Model cross
sections & corre



From D. Neudecker, S. Gundacker, H. Leeb *et al.*, ND2010, Jeju Isl., Korea

Definition of (Nuclear) Data Evaluation

A properly weighted combination of selected experimental data (and modeling results if needed).

Bayesian approaches (may use prior knowledge):

- “Non-model” GLSQ fit (standards)
- Model prior + experimental data:
 - Deterministic: Model Prior (Sens) + GLSQ
 - Stochastic (MC): **UMC, TMC + UMC**
 - Hybrid: Model Prior (MC) + GLSQ

All stochastic methods may be used to produce samples for TMC



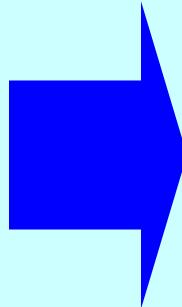
MONTE CARLO METHOD

D.L. Smith, “Covariance Matrices for Nuclear Cross-Sections
Derived from Nuclear Model Calculations”.
Report ANL/NDM-159, Argonne National Laboratory, 2005

$$\bar{\sigma}_i = \frac{1}{K} \sum_{k=1}^K \sigma_{ik} \quad V_{ij} = \overline{\sigma_i \sigma_j} - \overline{\sigma_i} \times \overline{\sigma_j} \quad i, j - \text{energy indexes}$$

Monte Carlo calculation of covariance first tested by A. Koning

Monte Carlo prior
+
GANDR (GLS)



A. Trkov and R. Capote, “Cross-Section Covariance Data”, Th-232 evaluation for ENDF/B-VII.0 (**MAT=9040 MF=1 MT=451**); Pa-231 and Pa-233 evaluations for ENDF/B-VII.0 (**MAT=9133 and 9137 MF=1 MT=451**), National Nuclear Data Center, BNL (<http://www.nndc.bnl.gov>), 15 December 2006.

D.W. Muir, **GANDR** project (IAEA),
Online at www-nds.iaea.org/gandr/.



UMC-G a.k.a.

UMC-B a.k.a.

“Garage” solution

“Breakfast” solution



VS



**D.L. Smith
San Diego 2007**

**R. Capote, A. Trkov
Port Jefferson 2008**



UMC-G (“Garage” Solution)

$\bar{\sigma}$ = collection of evaluated cross sections

\bar{y}_E = collection of experimental results. \bar{V}_E = corresponding cov. matrix

\bar{f} = collection of functions that relate $\bar{\sigma}$ to the data,
i.e., given $\bar{\sigma}$ we can calculate the equivalent to \bar{y}

$\bar{\sigma}_C$ = model calculated cross sections \bar{V}_C = corresponding cov. matrix

∴ $p(\bar{\sigma} | E, C)$ = probability density function for $\bar{\sigma}$ given experimental data “E” and calculated model-calculated prior results “C”

$$p(\bar{\sigma} | E, C) = C \exp \left\{ \left(-\frac{1}{2} \right) [\bar{y}_E - \bar{f}(\bar{\sigma})]^T \bar{V}_E^{-1} [\bar{y}_E - \bar{f}(\bar{\sigma})] + \left(-\frac{1}{2} \right) (\bar{\sigma} - \bar{\sigma}_C)^T \bar{V}_C^{-1} (\bar{\sigma} - \bar{\sigma}_C) \right\}$$

$$\bar{\sigma} = \sigma_1, \sigma_2, \dots, \sigma_i, \dots, \sigma_N$$

Donald L. Smith, ANL
Argonne National Laboratory
WPEC 2007 – SG24



UNIFIED MONTE CARLO (UMC-G)

D.L. Smith, “A Unified Monte Carlo Approach to Fast Neutron Cross Section Data Evaluation,” *Proceedings of the 8th International Topical Meeting on Nuclear Applications and Utilization of Accelerators*, Pocatello, July 29 – August 2, 2007, p. 736.

BAYES THEOREM & PRINCIPLE OF MAXIMUM ENTROPY

$$p(\sigma) = C \times \mathcal{L}(y_E, V_E | \sigma) \times p_0(\sigma | \sigma_C, V_C)$$

$$p_0(\sigma | \sigma_C, V_C) \sim \exp\{-\frac{1}{2}[(\sigma - \sigma_C)^T \cdot (V_C)^{-1} \cdot (\sigma - \sigma_C)]\}$$

$$\mathcal{L}(y_E, V_E | \sigma) \sim \exp\{-\frac{1}{2}[(y - y_E)^T \cdot (V_E)^{-1} \cdot (y - y_E)]\}, y=f(\sigma)$$

y_E, V_E : measured quantities with “n” elements

y_C, V_C : calculated using models with “m” elements

UMC based on $p(\sigma)$, GLS on the peak of the distribution



Unified Monte Carlo (UMC-B)

- 1) MC modeling (EMPIRE, TALYS, CCONE, CoH,...) $\{\sigma_i\}$
- 2) For each random set $\{\sigma_i\}$ we calculate $w^{\text{exp}}(\vec{\sigma}_i) = \mathcal{L}(\mathbf{y}_E, \mathbf{V}_E | \sigma_i)$

$$\mathcal{L}(\mathbf{y}_E, \mathbf{V}_E | \sigma_i) = \exp\left\{-\left(\frac{1}{2}\right)[(f(\sigma_i) - \mathbf{y}_E)^T \bullet (\mathbf{V}_E)^{-1} \bullet (f(\sigma_i) - \mathbf{y}_E)]\right\}$$

$$\langle \vec{\sigma} \rangle = \frac{\sum_{i=1}^N w^{\text{exp}}(\vec{\sigma}_i) \vec{\sigma}_i}{\sum_{i=1}^N w^{\text{exp}}(\vec{\sigma}_i)}, \quad \text{cov}(\vec{\sigma}_i, \vec{\sigma}_j) = \langle \vec{\sigma}_i \vec{\sigma}_j \rangle$$

OUTPUT: 1) $\langle \vec{\sigma} \rangle$, $\text{cov}(\vec{\sigma}_i, \vec{\sigma}_j) = \langle \vec{\sigma}_i \vec{\sigma}_j \rangle$

2) Stochastic set $\{\sigma_i\}$

RC, D. L. Smith, A. Trkov, M. Meghzifene, A New Formulation of the Unified Monte Carlo Approach (UMC-B) and Cross-Section Evaluation for ...,
J. ASTM International **9**, JAI104115 (2012)

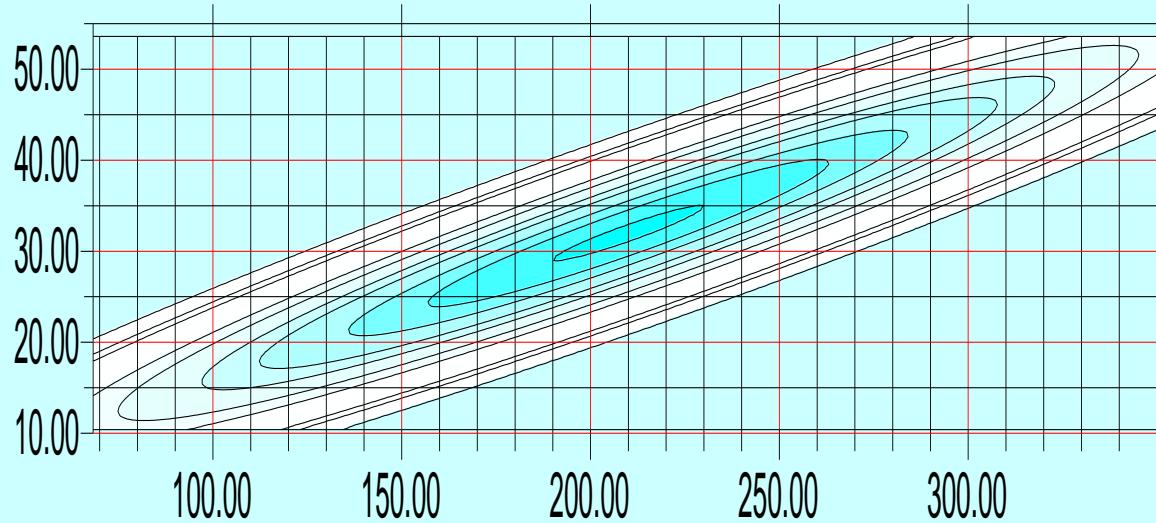


Two variable toy model: ratio experimental data



RATIO CASE

MODEL



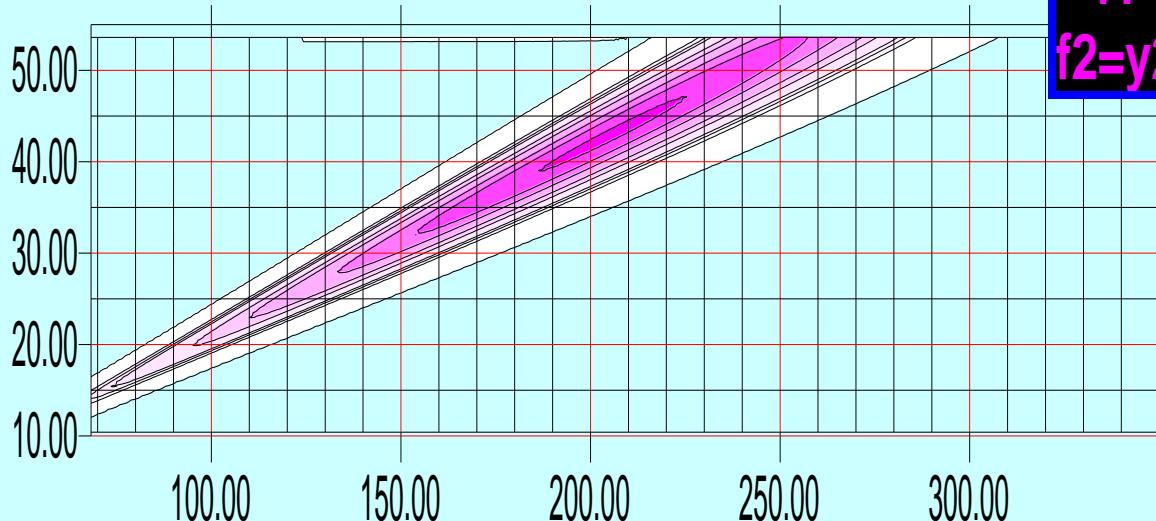
MODEL

$$y_1 = 210 \pm 63 \text{ (30\%)}$$

$$y_2 = 32 \pm 9.6 \text{ (30\%)}$$

$$\text{Cov}(1,2) = 0.95$$

EXPERIM



EXPERIM

$$f_1 = y_1 = 205.6 \pm 61.7 \text{ (30\%)}$$

$$f_2 = y_2/y_1 = 0.209 \pm 0.010 \text{ (5\%)} \sim 43$$

$$\text{Cov}(1,2) = 0.$$

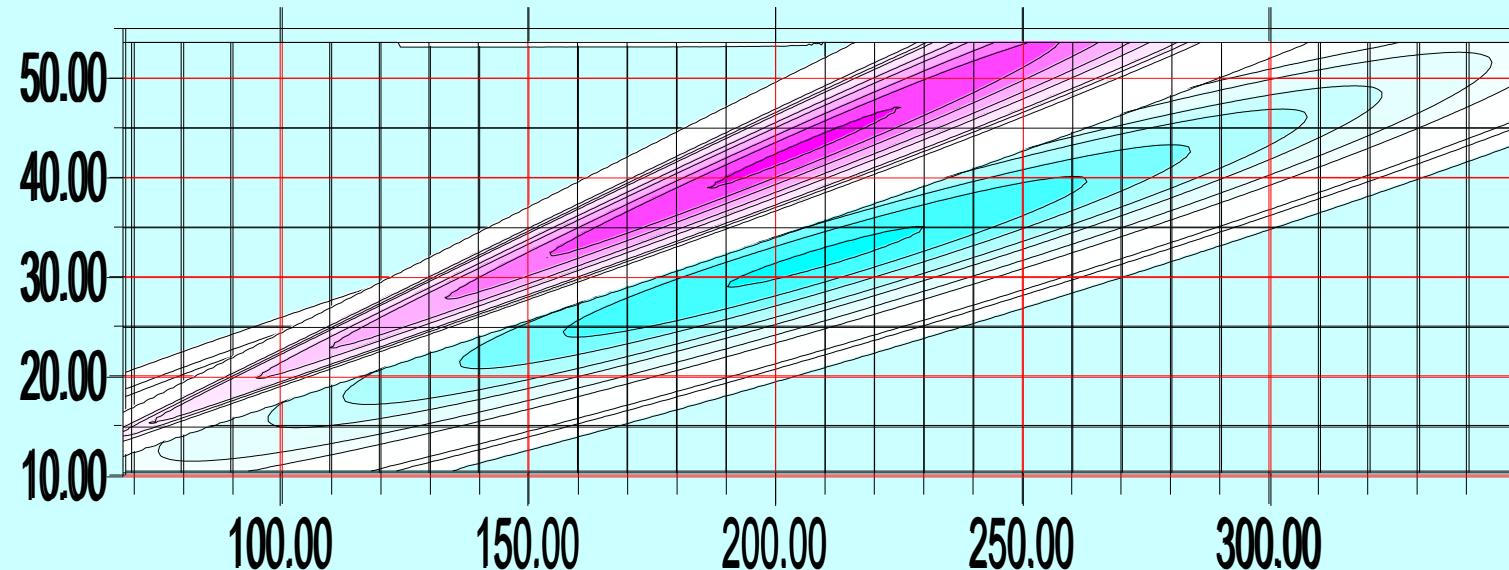
“y2”=43+/-2

vs 32+/- 9

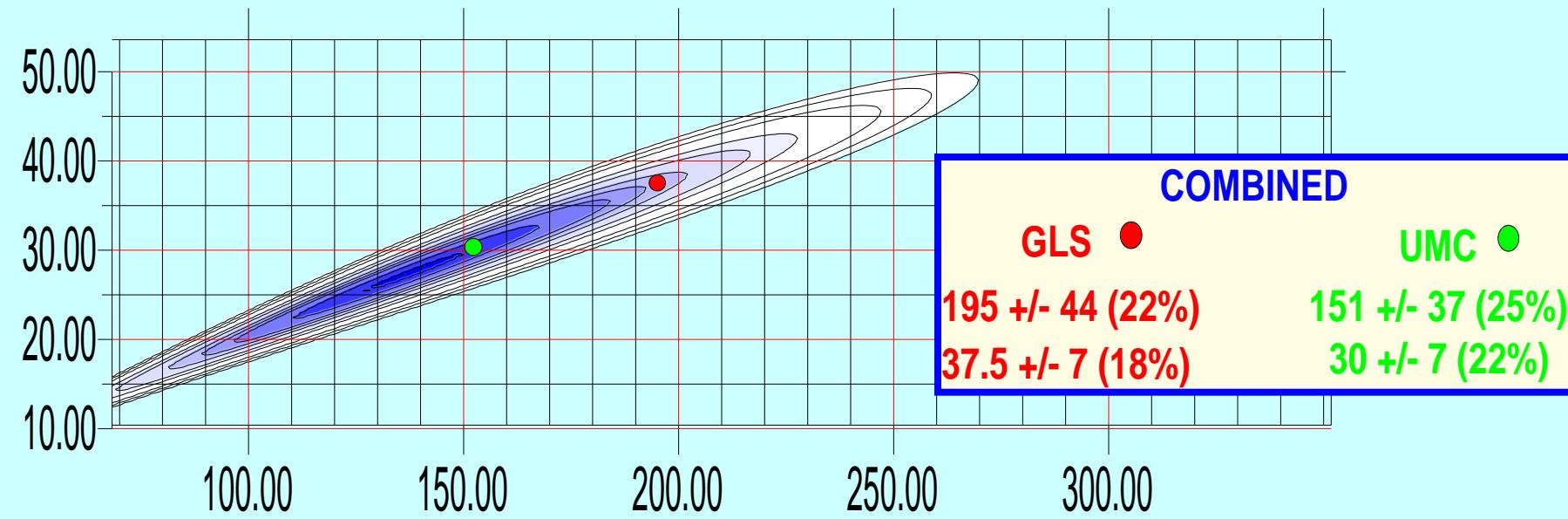


5% exp. ratio unc., 95% model correl.

EXPERIM



COMBINED



GLS FAILURE: ANALYSIS

5% exp. ratio unc.

**95% model correlation
(discrepant model vs data)**

<u>Quantity</u>	<u>Node 1</u>	<u>Node 2</u>	<u>Ratio</u>
BF/GLS	0.7767	0.7929	1.0209
METR/GLS	0.7728	0.7891	1.0210
METR/BF	0.9950	0.9951	1.0001

5% exp. ratio unc.

no model correlation

<u>Quantity</u>	<u>Node 1</u>	<u>Node 2</u>	<u>Ratio</u>
BF/GLS	1.0180	0.9795	0.9622
METR/GLS	1.0232	0.9850	0.9626
METR/BF	1.0051	1.0056	1.0004

30% exp. ratio unc.

95% model correlation

<u>Quantity</u>	<u>Node 1</u>	<u>Node 2</u>	<u>Ratio</u>
BF/GLS	1.0002	1.0007	1.0005
METR/GLS	0.9995	0.9998	1.0004
METR/BF	0.9992	0.9991	0.9999

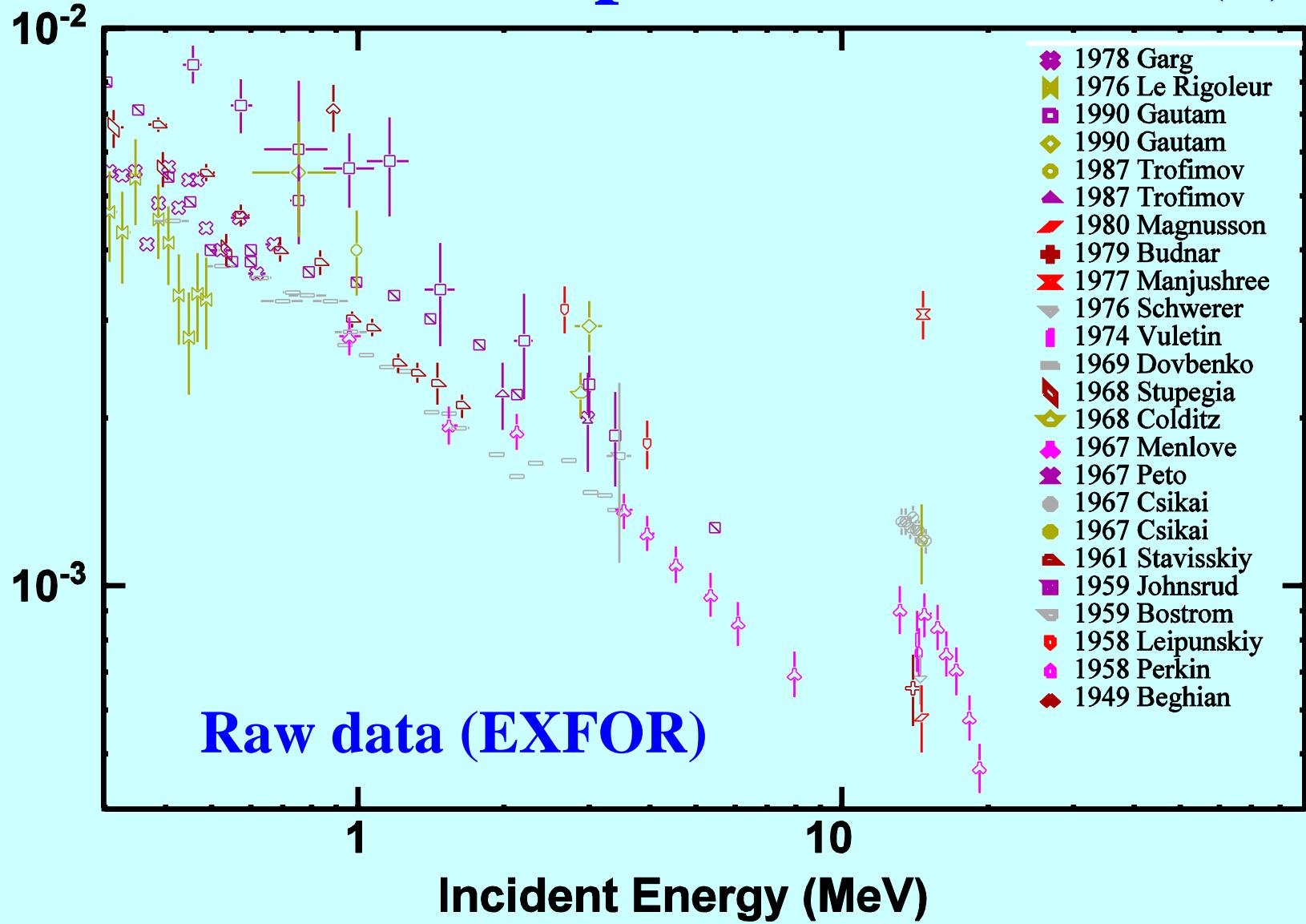


Evaluation of $^{55}\text{Mn}(\text{n},\gamma)$ cross section

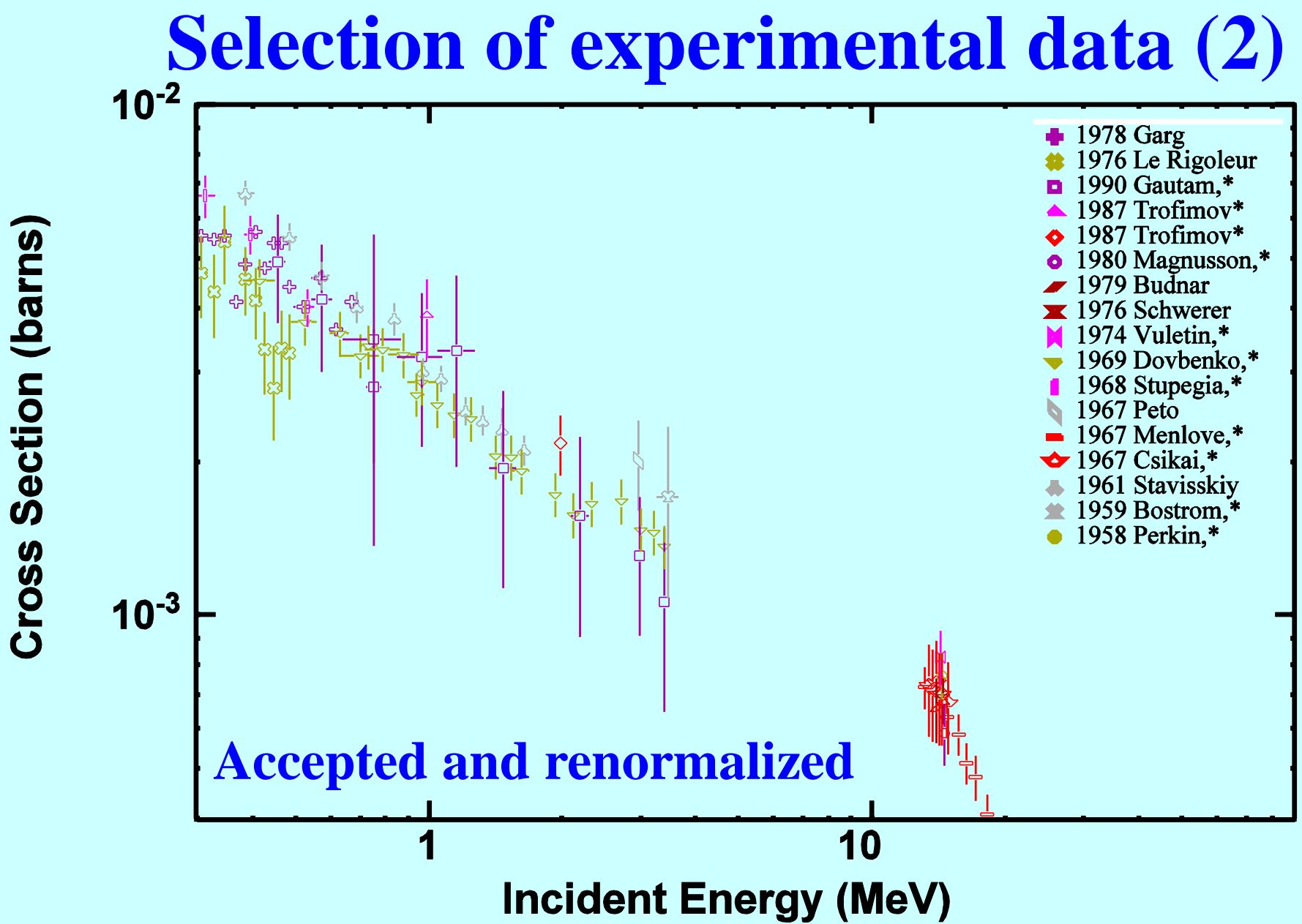


Selection of experimental data (1)

Cross Section (barns)



Selection of experimental data (2)



RIPL – Reference Input Parameter Library for Calculation of Nuclear Reactions and Nuclear Data Evaluations

R. Capote,¹ M. Herman,^{1,2} P. Obložinský,^{1,2} P.G. Young,³ S. Goriely,⁴ T. Belgya,⁵ A.V. Ignatyuk,⁶ A.J. Koning,⁷ S. Hilaire,⁸ V.A. Plujko,⁹ M. Avrigeanu,¹⁰ O. Bersillon,⁸ M.B. Chadwick,³ T. Fukahori,¹¹ Zhigang Ge,¹² Yinlu Han,¹² S. Kailas,¹³ J. Kopecky,¹⁴ V.M. Maslov,¹⁵ G. Reffo,¹⁶ M. Sin,¹⁷ E.Sh. Soukhovitskii,¹⁵ and P. Talou³



Available online at www.sciencedirect.com



ScienceDirect

**Nuclear Data
Sheets**

www.elsevier.com/locate/nds

Nuclear Data Sheets 108 (2009) 2655

$$\rightarrow p_0(\sigma \mid \sigma_c, V_c)$$

EMPIRE: Nuclear Reaction Model Code System for Data Evaluation

M. Herman^{1,*}, R. Capote², B.V. Carlson³, P. Obložinský¹, M. Sin⁴, A. Trkov⁵, H. Wienke⁶, and V. Zerkin²

¹ National Nuclear Data Center, Brookhaven National Laboratory, Upton, NY 11973-5000, USA

² Nuclear Data Section, International Atomic Energy Agency, Wagramer Strasse, A-1400 Vienna, Austria

³ Departamento de Física, Instituto Tecnológico de Aeronáutica, 12228-900, SP, São José dos Campos, Brazil

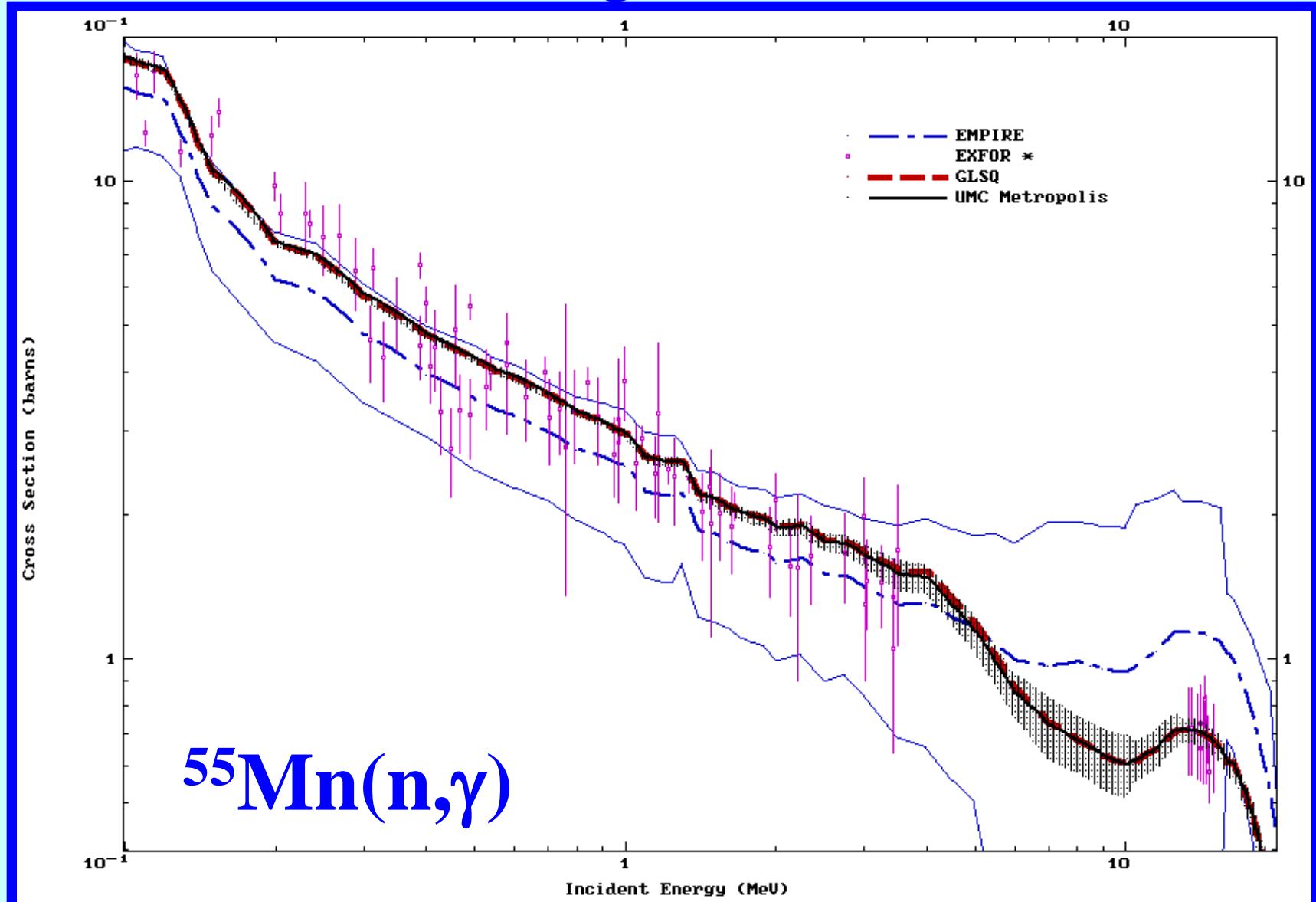
⁴ Nuclear Physics Department, Bucharest University, P.O. Box MG-11, Bucharest-Magurele, Romania

⁵ Jozef Stefan Institute, Reactor Physics Division R-1, Jamova 39, 1000 Ljubljana, Slovenia and

⁶ Belgonucleaire, Dessel, B2480, Belgium



UMC vs GLSQ: a real evaluation



Concluding remarks

- A careful selection and adjustment of raw experimental data **is a critical step** to achieve a consistent (non discrepant) database for the evaluation

- Linear case studied (no ratio data) : UMC-G and GLSQ are in excellent agreement

- Non-gaussian case studied (ratio data): GLSQ fails, UMC (Metr) solution obtained

- UMC-G (Metr), UMC-B and GLSQ applied to $^{55}\text{Mn}(\text{n},\gamma)$ case (energy range 100 keV to 20 MeV)





UMC sampling schemes

Brute Force approach: A set of independent $\{\sigma\}$

$$\bar{\sigma}_{Ck} - \psi [(\mathbf{V}_C)_{ii}]^{1/2} \leq \sigma_{ik} \leq \bar{\sigma}_{Ck} + \psi [(\mathbf{V}_C)_{ii}]^{1/2}$$

$$\sigma_{ik} = \bar{\sigma}_{Ck} + (2\gamma - 1) \psi [(\mathbf{V}_C)_{ii}]^{1/2}$$

METROPOLIS¹ approach: An stochastic Markov chain $\{\sigma\}$ distributed following $p(\sigma)$

$$\sigma' = \sigma(t) + (2\gamma - 1) \delta [(\mathbf{V}_C)_{ii}]^{1/2}, \text{ being } \sigma(t=0) = \bar{\sigma}_C$$

If $p(\sigma') > \gamma p(\sigma(t))$ then $\sigma(t+1) = \sigma'$; else $\sigma(t+1) = \sigma(t)$

RC and D.L. Smith, *Nucl. Data Sheets* **109**, 2768 (2008)

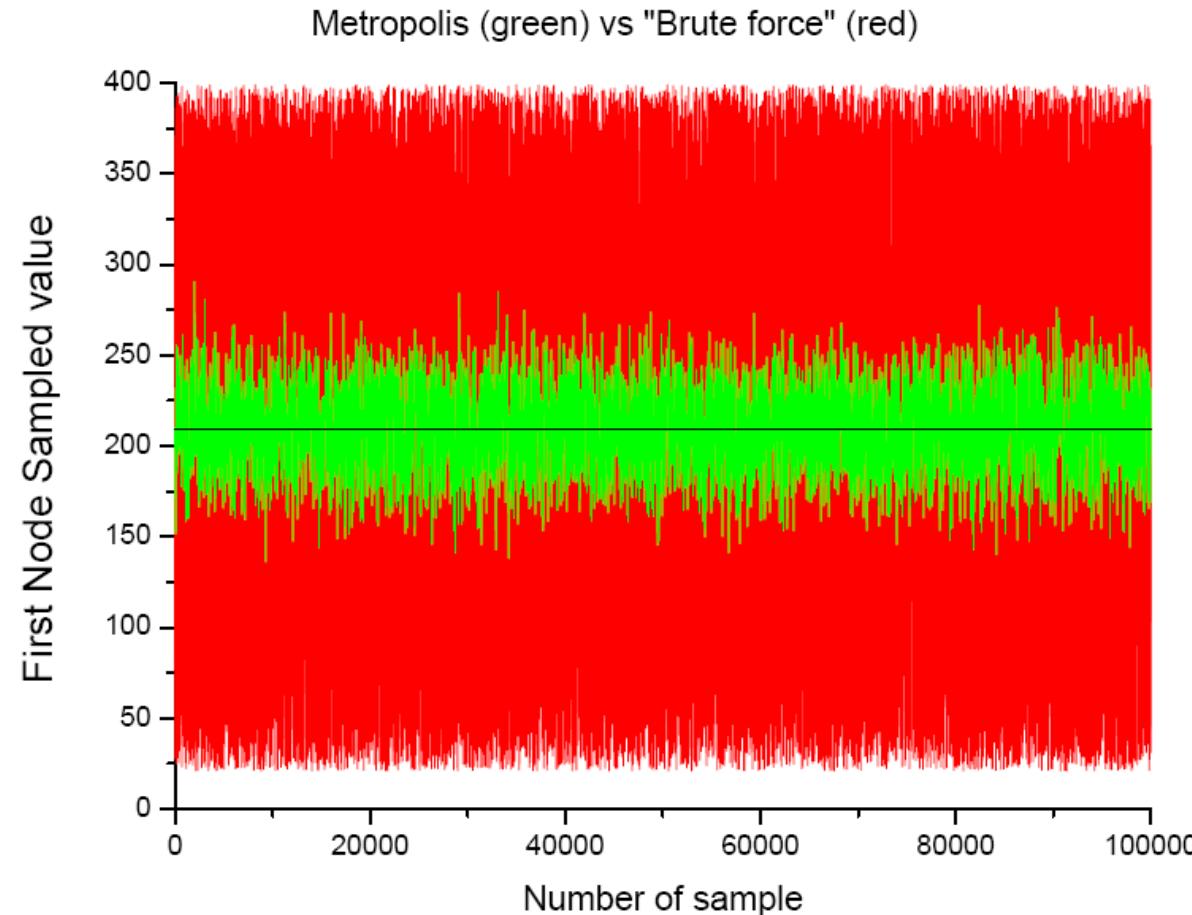
(1)N. Metropolis *et al.*, *J. Chem. Phys.* **21**, 1087 (1953)



UMC-G: BF vs Metropolis

Nuclear Data Sheets **109** (2008) 2768

An Investigation of the Performance of the Unified Monte Carlo Method of Neutron Cross Section Data Evaluation



UMC convergence

